Adaptive filtering by convex optimization

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Problem statement

We look to recover an unknown signal $x \in \mathbb{C}^T$, \mathcal{T} being a regular grid in \mathbb{R}^d , given noisy observations

$$y_{\tau} = x_{\tau} + \sigma \xi_{\tau}, \ \tau \in \mathcal{T}, \tag{1}$$

where ξ is the (complex-valued) white noise, $\xi_{\tau} \sim \mathcal{N}(0, \frac{1}{2}I_2)$.

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Optimal recovery:

Assume we want to estimate the value x_t of the signal at $t \in \mathcal{T}$.

Theorem [Ibragimov, Khas'minski, 1984, Donoho, 1994, etc, reform.] Let $\mathcal{X} \subset \mathbb{C}^{\mathcal{T}}$ be a convex compact and centrally symmetric set. Then for a variety of loss functions, the minimax, over $x \in \mathcal{X}$, risk of recovering x_t from noisy observations (1) is attained, within factor 1.2..., by a linear in y estimate, readily given along with its risk, by the solution to convex optimization problem [...]

Optimal recovery

In other words, if we are given a convex compact (and symmetric) set \mathcal{X} of signals (e.g., set of signals satisfying some regularity constraints) then a properly selected linear estimator

$$\mathbf{x}_t^* = \sum_{\tau \in \mathcal{T}} \varphi_\tau^* \mathbf{y}_\tau, \ \varphi^* \in \mathbb{C}^{\mathcal{T}},$$

is (quasi-) optimal on the class of all possible estimators.

• Computing the linear minimax estimator is "easy" for well-structured sets of signals (e.g., sets which can be described using CVX).

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Question:

Suppose that we do not know the class \mathcal{X} . Is it possible to "mimic" the oracle linear estimator φ^* , i.e. to construct an adaptive estimator (which only use observations) of comparable accuracy?

Problem reformulation

For the sake of simplicity, consider 1d situation, where the signal to recover $x \in \mathbb{C}^{\mathbb{Z}}$, and we are given n = 4T + 1 observations

$$y_{\tau} = x_{\tau} + \sigma \xi_{\tau}, \quad -2T < \tau < 2T, \tag{2}$$

Our objective may be either

- filtering estimation of x_{2T} (or x_{-2T}),
- interpolation estimation of x_t , -2T < t < 2T (e.g., x_0)
- prediction estimation of x_{2T+k}, (or x_{-2T-k}) for some k ∈ N₊.

We assume that the oracle estimator φ^* has bounded support – can be represented as a "linear filter" of length $\leq T + 1$. For instance, when estimating x_t , $-T/2 \leq t \leq T/2$,

$$x_t^* = \sum_{\tau = -T/2}^{T/2} \varphi_{\tau}^* y_{t-\tau} = [\varphi^* * y]_t.$$

Basic assumption

For the sake of simplicity, let us assume that we want to estimate x_0 . We say that $x \in \mathbb{C}_{-T}^{-T}$ if x vanishes outside the interval [-T, T].

We say that signal x is simple at t = 0 if there exists a (oracle) filter $\varphi^* \in \mathbb{C}_{-T/2}^{T/2}$, satisfying

- (small variance condition) $\|\varphi^*\|_2 \leq \frac{\rho}{\sqrt{T}}$,
- (small bias condition) for some $\theta > 0$ and all $-\frac{3T}{2} \le \tau \le \frac{3T}{2}$,

$$|x_{ au} - [\varphi^* * x]_{ au}| \leq rac{ heta \sigma
ho}{\sqrt{T}}.$$

More generally, for x which is simple at t, there exists φ^* of length T and a neighborhood of size O(T) of t where $\varphi^* * x$ reproduces x with "small bias".

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$$|x_{\tau} - [\varphi^* * x]_{\tau}| \leq \frac{\theta \sigma \rho}{\sqrt{T}}.$$

More generally, for x which is simple at t, there exists φ^* of length T and a neighborhood of size O(T) of t where $\varphi^* * x$ reproduces x with "small bias".

As a result, a simple at t = 0 signal x can be "well recovered" from y unformly over $-\frac{3T}{2} \le \tau \le \frac{3T}{2}$: $\mathbf{E}|x_{\tau} - [\varphi^* * (x + \sigma\xi)]_{\tau}|^2 = \sigma^2 \mathbf{E}|[\varphi^* * \xi]_{\tau}|^2 + |x_{\tau} - [\varphi^* * x]_{\tau}|^2$ $= \frac{\sigma^2 \rho^2}{T} + \frac{\theta^2 \sigma^2 \rho^2}{T} = O(1)\frac{\sigma^2 \rho^2}{T}.$

Classical example

Consider the problem of estimating a smooth function $f: [0,1] \to \mathbb{R}$ from noisy observations

$$y_i = f(i/n) + \sigma \xi_i, \quad i = 1, ..., n, \ \xi \sim \mathcal{N}(0, I_n).$$

The classical kernel estimator \hat{f}_t of f(t) with bandwidth h is

$$\widehat{f}(t) = \sum_{i=1}^{n} \frac{1}{2nh} K\left(\frac{t-i/n}{h}\right) y_i,$$

and K(t) : $[-1,1] \rightarrow \mathbb{R}$ is a kernel such that

$$\int_{-1}^{1} K(t) dt = 1, \ \ \int_{-1}^{1} K^{2}(t) dt =
ho^{2} < \infty.$$

Let $x_{\tau} = f(\tau/n)$, $\tau = 1, ..., n$, and let T = [2nh]. Then, the kernel estimator above can be rewritten for $T/2 + 1 \le t \le n - T/2$ as

$$\widehat{x}_t = \widehat{f}(t/n) = (\phi * y)_t, \ \phi_\tau = \frac{1}{T} \mathcal{K}\left(\frac{\tau}{T/2}\right), \ \tau = -T/2, ..., T/2.$$

Note that the ℓ_2 -norm of ϕ satisfies $\|\phi\|_2 \sim \rho/\sqrt{T}$, and if the kernel K and the bandwidth h are "properly chosen", the bias of the estimator is also $O(1)\rho/\sqrt{T}$.

Less classical example

Suppose that $f : [0,1] \to \mathbb{C}$ can be locally, when $x - h \le x \le x + h$, well-approximated by an exponential polynomial:

$$p(x) = \sum_{k=1}^{K} c_k x^{r_k} e^{i\omega_k x}$$

ł

with fixed frequencies $\omega_k \in \mathbb{C}$.



An exponential polynomial, K = 2

Note that for any T = 2nh > 2K there exists a kernel K_h^* , depending on the frequencies ω_k , of the norm $O_K(1)/\sqrt{T}$ which exactly reproduces p.

Less classical examples

When applied in the problem of estimation of f, kernel K_h^* , with properly chosen h, recovers f(x) with the "parametric rate" [J., Nemirovski, 2009, 2013]

$$O_{\kappa}(1)rac{\sigma^2}{nh}=O_{\kappa}(1)rac{\sigma^2}{T}.$$

Furthermore,

 The class of simple signals is quite rich, it contains, for instance, signals x_τ ∈ C which are close to solutions to homogeneous difference equations:

$$\sum_{k=1}^{K} w_k x_{\tau-k} = 0, \ w \in \mathbb{C}^{K}.$$

- This class allows for a calculus: linear combinations, modulations, liftings, "tensor products" of simple signals are also simple.
- More examples in multi-dimensional case [J., Nemirovski, 2009] ...

Problem reformulation

Question:

under these conditions, is it possible to design an "adaptive estimation" $\widehat{x}_0 = [\widehat{\varphi} * y]_0$ of x_0 which only relies upon observations $y \in \mathbb{C}_{-2T}^{2T}$, and such that

$$\left[{f E} |\widehat{x}_0 - x_0|^2
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Theorem 1 [lower bound].

For any $\rho \geq 1$, positive σ and $T \in \mathbb{N}$ large enough, one can point out a family \mathcal{F}_{ρ}^{T} of real signals on [-2T, 2T] such that

• for each signal $s \in \mathcal{F}_{\rho}^{\mathcal{T}}$ there exists a filter $\varphi^* \in \mathbb{R}_{-T/2}^{T/2}$ with $\|\varphi^*\|_2 = \frac{\rho}{\sqrt{T+1}}$, such that

$$\max_{-3T/2 \le \tau \le 3T/2} \left[\mathsf{E}((\varphi^* * y)_{\tau} - x_{\tau})^2 \right]^{1/2} = \frac{\sigma \rho}{\sqrt{T+1}};$$

• there is $c_0 > 0$ such that for any estimate \widehat{x}_0 of x_0 from observations (1) it holds

$$\sup_{x\in\mathcal{F}_{\rho}^{\mathcal{T}}}\left[\mathbf{E}(\widehat{x}_{0}-x_{0})^{2}\right]^{1/2}\geq c_{0}\frac{\sigma\rho}{\sqrt{T+1}}\;\rho\sqrt{\log(T+1)}\,.$$

Theorem 2 [upper bound].

Assume that x is simple at zero with known parameters ρ and θ . Then there is an estimate $\hat{x}_0(y)$ of x_0 such that

$$egin{array}{lll} \left[{\sf E} \left| \widehat{x}_0(y) - x_0
ight|^2
ight]^{1/2} &\leq & c \; rac{\sigma
ho}{\sqrt{ au}} \left[heta + \sqrt{\log(au+1)}
ight] \;
ho^2 \, . \end{array}$$

Furthermore, one has with probability 1-arepsilon ,

$$|\widehat{x}_0(y) - x_0| \leq c \frac{\sigma \rho}{\sqrt{T}} \left[\theta + \sqrt{\log\left(\frac{T+1}{\varepsilon}\right)} \right] \rho^2.$$

Naive approach – Empirical Risk minimization: For a signal $x \in \mathbb{C}^{\mathbb{Z}}$, $L \in \mathbb{N}_+$, and $1 \leq p \leq \infty$, let us denote

$$||x||_{L,p} = ||[x]_{-L}^{L}||_{p}$$

Define $\widehat{\varphi}$ as an optimal solution to

$$\min_{\varphi\in\mathbb{C}^{T+1}}\left\{\|y-\varphi\ast y\|^2_{3T/2,2}: \|\varphi\|_2\leq \frac{\rho}{\sqrt{T}}\right\}.$$

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$$\min_{\varphi \in \mathbb{C}^{T+1}} \left\{ \|y - \varphi * y\|_{3T/2,2}^2 : \|\varphi\|_2 \leq \frac{\rho}{\sqrt{T}} \right\}.$$

Note that φ^* is feasible, so that

$$\|y - \widehat{\varphi} * y\|_{3T/2,2}^2 \le \|y - \varphi^* * y\|_{3T/2,2}^2 = O_P(1) + \sigma^2 \|\xi\|_{3T/2,2}^2.$$

Therefore,

$$\begin{aligned} \|x - \widehat{\varphi} * y\|_{3T/2,2}^2 &= \|y - \widehat{\varphi} * y\|_{3T/2,2}^2 - \sigma^2 \|\xi\|_{3T/2,2}^2 - 2\sigma \langle \xi, x - \widehat{\varphi} * y \rangle_{3T/2} \\ &= O_P(1) + \underbrace{2\sigma^2 \langle \xi, \widehat{\varphi} * \xi \rangle_{3T/2}}_{O_P(\sqrt{T})} - 2\sigma \langle \xi, x - \widehat{\varphi} * x \rangle_{3T/2}. \end{aligned}$$

For $x \in \mathbb{C}^{\mathbb{Z}}$, let $F_T(x)$ be the Discrete Fourier Transform (DFT) of $[x]_{-T}^T$. We denote $||x||_{T,p}^* = ||F_Tx||_p$.

Lemma

Suppose that $\varphi^* \in \mathbb{C}_{-T/2}^{T/2}$ satisfies $\|\varphi^*\|_2 \leq \frac{\rho}{\sqrt{T}}$. Let also

$$\psi^* := (\varphi^* * \varphi^*) \in \mathbb{C}_{-T}^T.$$

Then ψ^* it holds

• $\|\psi^*\|_2 = \|\psi^*\|_{T,2}^* \le \|\psi^*\|_{T,1}^* \le \frac{\sqrt{2}\rho^2}{\sqrt{T}};$

• moreover, if x is simple at 0 then for $\tau : -T \le \tau \le T$, $|x_{\tau} - [\psi^* * x]_{\tau}| \le \frac{2\sigma\theta\rho^2}{\sqrt{\tau}}$.

Let $\widehat{\psi} \in \mathbb{C}_{-T}^{T}$ be an optimal solution of the following problem:

$$\min_{\psi \in \mathbb{C}_{-\tau}^{T}} \left\{ \| y - \psi * y \|_{\tau,2} : \| \psi \|_{\tau,1}^{*} \le \frac{\sqrt{2}\rho^{2}}{\sqrt{T}} \right\}.$$
 (P₁)

Then, as before, by the feasibility of ψ^{\ast}

$$\|\mathbf{y}-\widehat{\psi}*\mathbf{y}\|_{T,2} \leq \|\mathbf{y}-\psi^**\mathbf{y}\|_{T,2}.$$

• We have now better control of the cross-term $\langle \xi, \hat{\psi} * \xi \rangle_{\mathcal{T}}$ ("almost" the max of a convex function over a convex polyhedron):

$$\langle \xi, \widehat{\psi} * \xi \rangle_{\mathcal{T}} \leq \max_{\|\psi\|_{1}^{*} \leq \varrho^{2} \sqrt{2/\mathcal{T}}} \quad \langle \xi, \psi * \xi \rangle_{\mathcal{T}} = O_{P}\left(\log \mathcal{T}\right).$$

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• ...

• We finally get

$$\begin{split} & \left[\mathsf{E} \| x - [\widehat{\psi} * y] \|_{T,2}^2 \right]^{1/2} \leq C \sigma \rho (1+\theta) \left[\rho \sqrt{\log T} \right] \\ & \left[\mathsf{E} \left| x_0 - [\widehat{\psi} * y]_0 \right|^2 \right]^{1/2} \leq \frac{C \sigma \rho (1+\theta)}{\sqrt{T}} \left[\rho^2 \sqrt{\log T} \right] \end{split}$$

and

A variant

Let $\widehat{\psi}$ be an optimal solution to

$$\min_{\psi \in \mathbb{C}^{2T+1}} \left\{ \|y - \psi * y\|_{T,\infty}^* : \|\psi\|_{T,1}^* \le \frac{\sqrt{2}\rho^2}{\sqrt{T}} \right\}$$
 (P₂)

Theorem 3 [upper bound]

Consider the estimation $\widehat{x}_0(y) = \left[\widehat{\psi} * y\right]_0$ of x_0 . Then

$$\mathbf{E}\left[\left|x_{0}(y)-\widehat{x}_{0}\right|^{2}\right]^{1/2} \leq c \frac{\sigma \rho}{\sqrt{T}}\left[\varrho^{3} \sqrt{\log[T]}+\theta\right],$$

and, with probability $1 - \varepsilon$,

$$|\widehat{x}_0(y) - x_0| \leq c \frac{\sigma
ho}{\sqrt{T}} \left[\varrho^3 \sqrt{\log[T/\varepsilon]} + \theta
ight].$$

A summary

• Let (x_{τ}) admit, for some T, the estimate $x_{\tau}^* = [\varphi^* * y]_{\tau}$ with "bandwidth" T (i.e., with $\varphi^* \in \mathbb{C}_{-T/2}^{T/2}$) such that

$$\max_{\tau:|\tau-t|\leq 3T/2} \mathsf{E}\left\{|x_{\tau} - x_{\tau}^{*}|^{2}\right\} \leq \kappa^{2} := \frac{\sigma^{2}\mu^{2}}{T+1}$$
(3)

for some known $\mu \geq 1$.

Our objective is, assuming that *T* and μ are known, to recover x_t from observations [y]^{t+2T}_{t=2T} nearly as well as if we were using our hypothetic estimate x^{*}_t.

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- Our objective is, assuming that *T* and μ are known, to recover x_t from observations [y]^{t+2T}_{t-2T} nearly as well as if we were using our hypothetic estimate x^{*}_t.
- By (3), |φ^{*}|₂ ≤ μ/(√T+1) and x is simple.
 When applying Theorem 2 or 3 with ρ = μ, θ = 1, we conclude that the MSE of recovery x̂_t = [ψ̂ * y]_t is bounded, respectively, by

$$\underbrace{O(1)\mu^2 \sqrt{\log(T)\kappa}}_{\text{when using }(P_1)} \text{ or } \underbrace{O(1)\mu^3 \sqrt{\log(T)\kappa}}_{\text{when using }(P_2)}$$

Adaptation to ρ and ${\it T}$

In "practical applications" values of the parameter ρ and of the bandwidth ${\cal T}$ are unknown.

• The algorithms can be modified to be adaptive with respect to *ρ*. For in instance,(*P*₂) can be replaced with the "norm minimization" problem

$$\min_{\psi,r} \left\{ r: \begin{array}{l} \|y - \psi * y\|_{T,\infty}^* \leq 2\sigma(1+r)\sqrt{\log[T+1]}, \\ \|\psi\|_{T,1}^* \leq r(2T+1)^{-1/2}. \end{array} \right\}$$
(P'_2)

Instead of constrained problems, we can consider their penalized versions. For instance, (P_1) can be replaced with

$$\min_{\psi} \left\{ \|y - \psi * y\|_{T,2}^2 + \varkappa \sigma^2 \sqrt{2T + 1} \|\psi\|_{T,1}^* \right\}.$$
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• To choose a proper *T* we can use Lepski's algorithm, which amounts to compare estimators computed for various values of *T*.

Operational summary

When applying the proposed approach to "practical" recovery of a signal or an image

- For each point *t* of the grid we need
 - 1. choose a set of bandwidths $\{T_0 = 0, T_1 = 1, T_2 = 2, ..., T_K = n\}$,
 - 2. for each bandwidth T_k compute an approximate solution $\widehat{\psi}_{T_k,t}$ to (P_1) (or $(P_2), (P'_2), ...$)
 - 3. compute estimations $\widehat{x}_t[\mathcal{T}_k] = [\widehat{\psi}_{\mathcal{T}_k,t} * y]_t$ and aggregate them using Lepski's algorithm.
- To reduce the numerical cost, instead of proceeding point-wise, one can use block-wise update of filters...

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One needs to solve repeatedly problems (P_1) of the kind (or alike)

$$Opt = \min_{\psi \in \mathbb{C}_{-T}^{-}} \left\{ f(\psi) = \| y - y * \psi \|_{T,p}^* : \| \psi \|_{T,1}^* \le r \right\}, \ r > 0, \ p \in \{2, \infty\}.$$
 (P*)

Choosing the optimization tool 1

Note that (P_*) can be rewritten as a bilinear saddle-point problem: indeed, its objective,

$$f(\psi) = \max_{u \in \mathbb{C}^{2T+1}} \left\{ \langle u, F_T(y - y * \psi) \rangle, \|u\|_q \leq 1 \right\},$$

where $\frac{1}{a} + \frac{1}{a} = 1$.

When denoting $z = F_T(\psi)$,

$$Opt = \min_{\psi \in \mathbb{C}^{2T+1}} \max_{u \in \mathbb{C}^{2T+1}} \left\{ \langle u, \mathcal{A}z \rangle + \langle u, b \rangle : \|u\|_q \le 1, \ \|z\|_1 \le r \right\}, \tag{P_*}$$

where $q \in \{1,2\}$, $b = F_T(y)$, and \mathcal{A} is as follows:

$$\mathcal{A} z = F_{T} \left[y * F_{T}^{-1}(z) \right]$$

= $F_{T} \left[F_{2T}^{-1} \left\{ F_{2T} \left[0_{T}; y; 0_{T} \right] \cdot * F_{2T} \left[0_{2T}; F_{T}^{-1}(z); 0_{2T} \right] \right\} \right]$

(here $[x; 0_T]$ stands for the concatenation with zero vector of length T and .* is the Hadamard element-wise product).

- (P_*) is a bilinear saddle-point problem with domains which are balls of either ℓ_2 or ℓ_2/ℓ_1 -norm.
- Problems should be solved to (relatively) low accuracy a solution \hat{z} of accuracy

$$\epsilon(\widehat{z}) := f(\widehat{z}) - \operatorname{Opt} \leq \frac{1}{4}\operatorname{Opt}$$

will be largely sufficient.

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Under the premise, proximal First Order algorithms appear to be methods of choice.

Proximal algorithms for bilinear saddle-point optimization

- $1/\epsilon$ complexity estimates (or even $1/\sqrt{\epsilon}$ under "favorable circumstances").
- Accuracy certificates are available "at no cost".
- Favorable geometry of the problem domain simple O(n) proximal computation.
- Fully profit from fast gradient computation $O(n \log n)$ cost per iteration.

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We have a choice of at least 2 efficient techniques:

- Extra-gradient algorithms for saddle-point problems (Mirror-Prox [Nemirovski, 2003], Dual Extrapolation [Nesterov, 2003], etc)
- Smoothing [Nesterov, 2003]: replace f(z) = max_{||u||q≤1}⟨u, Az⟩ with its "Nesterov's smoothing":

$$f_{\gamma}(z) = \max_{\|u\|_q \leq 1} \left\{ \langle u, \mathcal{A}z \rangle + \gamma \vartheta(u) \right\},\,$$

where ϑ is 1-strongly convex with respect to $\|\cdot\|_q$ -norm; then apply to f_{γ} Nesterov's accelerated algorithm for smooth optimization.

Comparing the contenders: theory

Nesterov accelerated algorithm:

- allows for easily implementable Euclidean and non-Euclidean prox and adaptive stepsize strategies;
- receives a "special mention" in the case of ℓ₂-norm minimization: instead of smoothing one can minimize the squared norm.
 In this case, accelerate algorithm exhibits 1/√ε complexity for ε ≪ Opt.
- allows for the easily implementable warm start: the theoretical accuracy estimate depends on the initial distance to the optimum (though not on the sub-optimality of the initial solution).
- However, smoothing implementation (in its "basic form") requires to fix from the start the regularisation parameter $\gamma \asymp 1/\epsilon$, what results in curbed convergence rates.

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- However, smoothing implementation (in its "basic form") requires to fix from the start the regularisation parameter $\gamma \asymp 1/\epsilon$, what results in curbed convergence rates.

Extra-gradient algorithms:

- allows for easily implementable Euclidean and non-Euclidean prox and adaptive stepsize strategies;
- can be seen as "online adjustment" of the regularization γ .
- On the other hand, no simple "warm start" strategy is available in this case.

Comparing the contenders: experiments



 ℓ_2 -norm minimization. Filter length T = 200, modulated 2nd order polynomial. Left plot – absolute error, right plot – relative error as a function of iteration count.

Simulation experiment: adaptive recovery



Comparison with Atomic Soft Thresholding (AST), a.k.a. spectral Lasso by [Bhaskar et al., 2013, Tang et al., 2013]

Modulated 4th order polynomial, SNR=1. AST over-sampling factor $\kappa = 4$. 23/27

Simulation experiment: adaptive recovery



Modulated 4nd order polynomial, SNR=1. AST over-sampling factor $\kappa = 4$.

Simulation experiments: sum of harmonic oscillations



Sum of 4 oscillations. AST over-sampling factor $\kappa = 4$.

Sum of harmonic oscillations: zoomed image



Filtering recovery, MSE=8.99Dasso recovery, MSE=66.5866



Sum of 4 oscillations. AST over-sampling factor $\kappa = 4$.

Simulation experiments: Brodatz picture



Brodatz D75 picture, SNR=1. AST over-sampling factor $\kappa = 4$. MISE_{Adapt}=3.2748e+03, MISE_{AST}=3.2514e+03.